Chap 5. Continuous V.U's.  
§ 5.1 Introduction.  
Def. X is called a Cont. r.v. (or absolute  
cont. r.u.) if there is a non-negative  
function defined on 
$$(-\infty, \infty)$$
 such that  
 $P\{X \in B\} = \int_B f(x) dx$   
for all "measurable" sets  $B \subset (-\infty, \infty)$ .  
Remark: The measurable sets include all intervals,  
and the countable unions / intersections of intervals.  
for  
 $A = \int_A^b f(x) dx$ 

 $P\{a \leq X \leq b\} = \int_{a}^{b} f(x) dx$ = Area of the shaded region. · We call f the prob. density function (pdf) of X. Example 2. Let X be a cts r.u. with pdf  $f(x) = \begin{cases} C(4x-2x^{2}) & \text{if } x \in (0,2) \\ 0 & \text{otherwise.} \end{cases}$ (1) Find the value of C (2) Find P{X < 1} •  $\int_{-\infty}^{\infty} f(x) dx = 1$ . Since  $P\{X \in (-\infty, \infty)\} = P\{S\} = I$ 

Solution:  $1 = \int_{-\infty}^{\infty} f(x) dx$  $= \int_{D}^{2} C(4x-2x^{2}) dx$  $= \left| \zeta \cdot \left( 2 \chi^2 - \frac{2}{3} \chi^3 \right) \right|_0^2$  $= C \cdot \left( \frac{\delta}{3} - \frac{2}{3} \times \delta \right) = \frac{\delta}{3} C$ Hence  $C = \frac{3}{8}$  $P\{X \le 1\} = \int_{-\infty}^{1} f(x) dx$  $= \int_{0}^{1} \frac{3}{8} (4x - 2x^{a}) dx$  $= \frac{3}{8} \left( 2\chi^2 - \frac{2}{3}\chi^3 \right)^{1/2}$ =  $\frac{1}{2}$ .

Exer. 1. Suppose X has a Pdf  $f(x) = \begin{cases} \lambda e^{\frac{-x}{100}} & \text{if } x > 0 \\ 0 & \text{otherwise} \end{cases}$ (1) Find the value of  $\lambda$ . (2) Find P{ X > 100} Solution:  $1 = \int_{-\infty}^{\infty} f(x) dx$  $= \int_{\infty}^{\infty} y e^{\frac{1}{2} y} dx$  $= \lambda \cdot \left( -100 \left( e^{-\frac{x}{100}} \right) \right)^{\infty}$  $= \lambda \cdot \iota \circ \circ$ Hence  $\lambda = \frac{1}{100}$ .

 $P\{X > 100\} = \int_{100}^{+\infty} f(x) dx$  $= \begin{pmatrix} 400 & -\frac{1}{100} \\ -\frac{1}{100} & -\frac{1}{$  $= -e^{-\frac{\chi}{100}} \int_{-\infty}^{\infty}$  $= e^{-1}$ 11 § 5.2 Expectation of a cts. r.v. Recall that the expectation of a discrete r.u. X is defined by  $E[X] = \sum x P\{X=x\},$ where the summation is taken over all the possible values that X can take on. We can not directly use this method to define the expectation of a cts r.u., since in the cts case, X will take on uncountably diff values, and more over, P{X=x}=o for all xER.

Def. Let X be a cts r.v. with pdf f. Then we define  $E[X] = \int_{\infty}^{\infty} x f(x) dx.$ Intuitive i dea : In the continuous core, we make a partition of  $(-\infty, \infty)$  by  $(X_n)_{n=-\infty}$  such that  $\chi_{n+1} - \chi_n = \Delta \chi.$ Then  $\sum x_n \cdot p \left\{ x_n < X \leq x_{n+1} \right\}$  $= \sum_{n} X_{n} \int_{X_{n}}^{X_{n}+4X} f(x) dx$  $\approx \sum_{n} x_{n} f(x_{n}) \cdot \Delta x \longrightarrow \int_{x_{n}}^{\infty} f(x) dx$  $\text{ as } \Delta x \rightarrow 0.$ 

Example 2. X is said to be uniformly distributed on [0,1] if it has the following density  $f(x) = \begin{cases} 1 & \text{if } x \in [0, 1] \\ 0 & \text{otherwise} \end{cases}$ Find E[X]. Solution:  $E[X] = \int_{-\infty}^{\infty} x f(x) dx$  $=\int_{D}^{1} x \cdot 1 dx$  $= \frac{\chi^{2}}{2} \begin{bmatrix} 1 \\ -2 \end{bmatrix} = \frac{1}{2}.$ Below we consider the expectation of functions of cts r. u.'s.

Prop 3. Let X be a cts n.v. with density f.  
Let 9: 
$$[R \rightarrow IR$$
. Then  
 $E[g(X)] = \int_{-\infty}^{\infty} g(x) f(x) dx$ .  
We prove the above prop only in the cone that  $9 \ge 0$ .  
To this end, we first prove the following.  
Lem 4. Let Y be a non-negative cts r.v. Then  
 $E[Y] = \int_{0}^{\infty} P\{Y > y\} dy$ .  
To prove Lem 4, we need to use the following result:  
 $\int_{a}^{b} (\int_{c}^{d} f(x, y) dx) dy = \int_{c}^{d} (\int_{a}^{b} f(x, y) dy) dx$   
When f is non-negative. This is a special version  
of the Fubini Theorem.

$$= \int_{0}^{\infty} f^{(x)} \left( \int_{0}^{x} 1 dy + \int_{x}^{\infty} 0 dy \right) dx$$

$$= \int_{0}^{\infty} x f(x) dx$$

$$= E[Y].$$

$$Proof of Prop. 3: By Lem 4,$$

$$E[g(X)] = \int_{0}^{\infty} P\{g(X) > y\} dy.$$

$$W_{n}'te \beta := \{x \in \mathbb{R} : g(x) > y\}. Thun$$

$$P\{g(X) > y\} = P\{x \in \mathbb{R}\} = \int_{\mathbb{R}} f_{(x)} dx = \int_{\{x : g(x) > y\}} f_{(x)} dx.$$
Hence
$$E[g(X)] = \int_{0}^{\infty} \int_{\{x : g(x) > y\}} f_{(x)} dx dy$$

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So by the Fubini Thm,  

$$E[g(x)] = \int_{-\infty}^{\infty} \left( \int_{0}^{\infty} h(x,y) f(x) dy \right) dx$$

$$= \int_{-\infty}^{\infty} f(x) \left( \int_{0}^{\infty} h(x,y) dy + \int_{0}^{\infty} h(x,y) dy \right)$$

$$= \int_{-\infty}^{\infty} f(x) \left( \int_{0}^{g(x)} 1 dy + \int_{0}^{\infty} h(x,y) dy \right) dx$$

$$= \int_{-\infty}^{\infty} f(x) \left( \int_{0}^{g(x)} 1 dy + \int_{0}^{\infty} (x) dy \right) dx$$

$$= \int_{-\infty}^{\infty} f(x) g(x) dx.$$
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Let X be a cts r.u. Define the Variance of X by  
Var(X) = E[(X-\mu)^{2}], where 
$$\mu$$
 = E[X].  
Prop. 1. Var(X) = E[X<sup>2</sup>] -  $\mu^{2}$ .  
Pf. Let f be the density of X. Then  
(Var(X) = E[(X-\mu)^{2}]  
=  $\int_{-\infty}^{\infty} (x^{2}-2x\mu + \mu^{2}) f(x) dx$   
=  $\int_{-\infty}^{\infty} x^{2} f(x) dx - 2\mu \int_{-\infty}^{\infty} x f(x) dx$   
 $+\mu^{2} \int_{-\infty}^{\infty} f(x) dx$   
=  $\int_{-\infty}^{\infty} x^{2} f(x) dx - 2\mu \cdot \mu + \mu^{2}$   
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=  $\int_{-\infty}^{\infty} x^{2} f(x) dx - 2\mu \cdot \mu + \mu^{2}$   
=  $\int_{-\infty}^{\infty} x^{2} f(x) dx - \mu^{2}$   
= E[X<sup>2</sup>] -  $\mu^{3}$ .

§ 5.3. Uniform distributions.  
Def. A cts X is said to be Uniformly distributed  
on [a, b] if it has a density  

$$f(x) = \begin{cases} \frac{1}{b-a} & \text{if } x \in (a,b) \\ 0 & \text{otherwise} \end{cases}$$
  
for  $a = \begin{cases} 0 & \text{otherwise} \end{cases}$   
We call X is a Unif. r.v on [a,b], or say  
X has a Unif distribution on [a,b].

Example 2. Calculate E[X] and V(X) for a unif r.v. X on [a, b]. Solution :  $E[X] = \int_{-\infty}^{\infty} x f(x) dx$  $= \int_{a}^{b} x \frac{1}{b-a} dx$  $= \frac{\pm x^{2}}{b-a} |_{a}^{b} = \frac{\pm (b^{2}-a^{2})}{b-a}$  $= \underline{atb}$ .  $E[X^{2}] = \int_{-\infty}^{\infty} x^{2} f(x) dx$  $= \int_{a}^{b} x^{2} \frac{1}{b-a} dx$  $= \frac{1}{3} \frac{b^3 - a^3}{(b - a)} = \frac{a^2 + ab + b^2}{3}$ 

Henre  $V(X) = \frac{a^2 + ab + b^2}{3} - \frac{(a + b)^2}{4}$  $= \frac{(a-b)^{2}}{12}$ · Def. (Cumulative distribution function) Let X be a cts r.v with density f. We define  $F_{X}(b) = P\{X \leq b\}$  $= \int_{-\infty}^{b} f(x) dx.$ 

Prop.3. If f is cts at b, then  $F_{\vee}(b) = f(b).$ Pf. Notice that for UER, Uto,  $\frac{F_{x}(b+u) - F_{x}(b)}{u} = \frac{\int_{-\infty}^{b+u} f(x) dx - \int_{-\infty}^{b} f(x) dx}{u}$  $= \frac{1}{u} \int_{L}^{btu} f(x) dx$ since f is cts at b, so f(x)-f(b) is close to o when x is close to b, hence as u >0, In Shufcxidx > f(b)  $\overline{M}$ .

Example 4. Let X be a unif r.v over (0,1). Find the density of X2. <u>Solution</u>:  $F_{\chi^2}(b) = P \{ \chi^2 \leq b \}$  $= \begin{cases} P\{-\sqrt{b} \le X \le \sqrt{b}\}, & \text{if } b > 0 \\ 0 & \text{otherwise} \end{cases}$  $= \begin{cases} 1 & \text{if } b > 1, \\ \sqrt{b} & \text{if } b \in (0, 1) \end{cases}$ if be(0,1), if b<0.

Taking derivative we obtain  $f_{\chi^2}(b) = \begin{cases} 1 & \text{if } b \in (0, 1), \\ 0 & \text{other wise.} \end{cases}$ 14