

Example 1. Let A, B be two **fixed** subsets of \mathbb{R} . Suppose that random variables X, Y have joint PDF

$$f(x, y) = \begin{cases} g(x)h(y) & \text{if } x \in A, y \in B \\ 0 & \text{otherwise.} \end{cases}$$

for some functions $g, h: \mathbb{R} \rightarrow \mathbb{R}$. Then X and Y are independent.

Proof. Let χ_C denote the indicator function for $C \subset \mathbb{R}^2$. Then

$$f(x, y) = g(x)h(y)\chi_{A \times B} = \left(g(x)\chi_{A \times Y}\right)\left(h(y)\chi_{X \times B}\right).$$

Hence X and Y are independent. **Alternatively**, if $x \notin A$, then $f_X(x) = 0$ and if $x \in A$

$$f_X(x) = \int_{\mathbb{R}} f(x, y)dy = \int_B g(x)h(y)dy = g(x) \int_B h(y)dy.$$

If $y \notin B$, then $f_Y(y) = 0$ and if $y \in B$

$$f_Y(y) = \int_{\mathbb{R}} f(x, y)dx = \int_A g(x)h(y)dx = h(y) \int_A g(x)dx.$$

By the unit integral of f , we have $\left(\int_A g(x)dx\right)\left(\int_B h(y)dy\right) = 1$. Then for $x \in A$ and $y \in B$,

$$f_X(x)f_Y(y) = g(x)h(y)\left(\int_A g(x)dx\right)\left(\int_B h(y)dy\right) = g(x)h(y) = f(x, y),$$

and for $x \notin A$ or $y \notin B$, $f(x, y) = 0 = f_X(x)f_Y(y)$. Hence X and Y are independent. \square

Recall

Independence of n random variables

For $n \geq 2$, let X_1, \dots, X_n be n random variables. The *joint cumulative distribution function* (joint CDF) of X_1, \dots, X_n is

$$F(a_1, \dots, a_n) := P\{X_1 \leq a_1, \dots, X_n \leq a_n\}, \quad \forall a_1, \dots, a_n \in \mathbb{R}.$$

- X_1, \dots, X_n are *jointly continuous* if there exists $f: \mathbb{R}^n \rightarrow [0, \infty)$ such that

$$P\{(X_1, \dots, X_n) \in C\} = \int \cdots \int_C f(x_1, \dots, x_n)dx_1 \cdots dx_n$$

for all “measurable” sets $C \subset \mathbb{R}^n$.

- X_1, \dots, X_n are *independent* if

$$P\{X_1 \in A_1, \dots, X_n \in A_n\} = P\{X_1 \in A_1\} \cdots P\{X_n \in A_n\}, \quad \forall A_1, \dots, A_n \subset \mathbb{R},$$

which is equivalently characterized by the joint CDF

$$F(a_1, \dots, a_n) = F_{X_1}(a_1) \cdots F_{X_n}(a_n), \quad \forall a_1, \dots, a_n \in \mathbb{R}.$$

Example 2 (Pairwise independence $\not\Rightarrow$ independence). Let $X \sim \text{Bern}(\frac{1}{2})$ and $Y \sim \text{Bern}(\frac{1}{2})$.

Suppose that X, Y are independent. Define $Z = \begin{cases} 1 & \text{if } X \neq Y \\ 0 & \text{if } X = Y. \end{cases}$

Then the joint PMF of X, Y, Z is

$$p(x, y, z) = \begin{cases} \frac{1}{4} & (x, y, z) = (0, 0, 0) \\ \frac{1}{4} & (x, y, z) = (0, 1, 1) \\ \frac{1}{4} & (x, y, z) = (1, 0, 1) \\ \frac{1}{4} & (x, y, z) = (1, 1, 0) \\ 0 & \text{otherwise.} \end{cases}$$

Hence

$$p_X(a) = p_Y(a) = p_Z(a) = \begin{cases} \frac{1}{2} & a = 1 \\ \frac{1}{2} & a = 0 \\ 0 & \text{otherwise.} \end{cases}$$

and

$$p_{X,Y}(a, b) = p_{X,Z}(a, b) = p_{Y,Z}(a, b) = \begin{cases} \frac{1}{4} & (a, b) = (0, 0) \\ \frac{1}{4} & (a, b) = (0, 1) \\ \frac{1}{4} & (a, b) = (1, 0) \\ \frac{1}{4} & (a, b) = (1, 1) \\ 0 & \text{otherwise.} \end{cases}$$

Thus X, Y, Z are pairwise independent. However, since

$$p(0, 0, 0) = \frac{1}{4} \quad \text{while} \quad p_X(0)p_Y(0)p_Z(0) = \frac{1}{8},$$

we have X, Y, Z are NOT independent.

Convolution formula for sum of independent random variables

Let X, Y be **independent** random variables. Then

$$\begin{cases} \text{if } X, Y \text{ joint continuous:} & f_{X+Y} = f_X * f_Y, \text{ that is } \forall z \in \mathbb{R}, f_{X+Y}(z) = \int_{-\infty}^{\infty} f_X(z-y)f_Y(y)dy. \\ \text{if } X, Y \text{ discrete:} & p_{X+Y} = p_X * p_Y, \text{ that is } \forall z \in \mathbb{R}, p_{X+Y}(z) = \sum_y p_X(z-y)p_Y(y). \end{cases}$$

In particular,

- If $X, Y \sim U(0, 1)$ independent, then $f_{X+Y}(z) = \begin{cases} z & 0 \leq z \leq 1 \\ 2 - z & 1 < z \leq 2 \\ 0 & \text{otherwise.} \end{cases}$
- If $X \sim N(\mu_1, \sigma_1^2), Y \sim N(\mu_2, \sigma_2^2)$ independent, then $X + Y \sim N(\mu_1 + \mu_2, \sigma_1^2 + \sigma_2^2)$. Combining with the previous result, we have for $a, b, c \in \mathbb{R}, aX + bY + c \sim N(a\mu_1 + b\mu_2 + c, a^2\sigma_1^2 + b^2\sigma_2^2)$.
- If $X \sim \text{Poisson}(\lambda_1), Y \sim \text{Poisson}(\lambda_2)$ independent, then $X + Y \sim \text{Poisson}(\lambda_1 + \lambda_2)$.

Example 3. For $m, n \in \mathbb{N}$ and $0 \leq p \leq 1$, let $X \sim \text{Bin}(n, p)$ and $Y \sim \text{Bin}(m, p)$ be independent. Then $X + Y \sim \text{Bin}(n + m, p)$.

Proof. Let $z \in \{0, \dots, n + m\}$. With convention $\binom{n}{k} = 0$ if $k > n$ and by the independence,

$$\begin{aligned} p_{X+Y}(z) &= p_X * p_Y(z) \\ &= \sum_{y=0}^z \binom{n}{z-y} p^{z-y} (1-p)^{n-(z-y)} \cdot \binom{m}{y} p^y (1-p)^{m-y} \\ &= \sum_{y=0}^z \binom{n}{z-y} \binom{m}{y} p^z (1-p)^{m+n-z} \\ &= \binom{m+n}{z} p^z (1-p)^{m+n-z} \end{aligned}$$

where the third equality is by $\sum_{y=0}^z \binom{n}{z-y} \binom{m}{y} = \binom{m+n}{z}$. When $z \notin \{0, \dots, n + m\}$, we have $p_{X+Y}(z) = 0$. Hence $X + Y \sim \text{Bin}(n + m, p)$. \square