Recall

Common discrete distributions

Bernoulli r.v.. with parameter $p \in [0, 1]$. Denote $X \sim Bern(p)$.

- (1) X is the outcome of a trial that succeeds with probability p and fails with probability 1 p.
- (2) The PMF of X is $p(1) = P\{X = 1\} = p$ and $p(0) = P\{X = 0\} = 1 p$.

Note E[X] = p and Var(X) = p(1-p).

Binomial r.v.. with parameter (n, p). Denote $X \sim Bin(n, p)$.

- (1) X is the number of successes that occur in the n independent Bernoulli trials with parameter p.
- (2) The PMF of X is $p(k) = P\{X = k\} = \sum_{k=0}^{n} {n \choose k} p^k (1-p)^{n-k}$ for $k = 0, \dots, n$.
- (3) Let X_1, \ldots, X_n be independent Bernoulli r.v.s with parameter p. Then $X = \sum_{k=1}^n X_k$.

Note E[X] = np and Var(X) = np(1-p).

Poisson r.v.. with parameter $\lambda > 0$. Denote $X \sim Poisson(\lambda)$.

(1) The PMF of X is
$$p(k) = P\{X = k\} = \frac{e^{-\lambda}\lambda^k}{k!}$$
 for $k = 0, 1, 2, ...$

Note $E[X] = \operatorname{Var}(X) = \lambda$.

Geometric random variable & computing examples

Example 1 (Geometric r.v. with parameter p). Denote $X \sim Geom(p)$.

- (1) X is the number of independent Bernoulli trials with parameter $p \in (0, 1)$ such that first success occur.
- (2) By independence, the PMF is $p(k) = (1-p)^{k-1}p$ for k = 1, 2, 3, ...

Then we show $E[X] = \frac{1}{p}$ and $Var(X) = \frac{1-p}{p^2}$.

By definition, $E[X] = \sum_{k=1}^{\infty} k(1-p)^{k-1}p$. Note that for $x \in (-1,1)$, (by a result about the uniform convergence of power series) we have

$$\sum_{k=1}^{\infty} kx^{k-1} = \sum_{k=0}^{\infty} (x^k)' = (\sum_{k=0}^{\infty} x^k)' = (\frac{1}{1-x})' = \frac{1}{(1-x)^2}.$$

Then setting x = 1 - p leads to $E[X] = p \sum_{k=1}^{\infty} k(1-p)^{k-1} = p \frac{1}{[1-(1-p)]^2} = \frac{1}{p}.$

To obtain Var(X), it suffices to compute $E[X^2]$

$$\begin{split} E[X^2] &= \sum_{k=1}^{\infty} k^2 (1-p)^{k-1} p \\ &= \sum_{k=1}^{\infty} (k-1+1)^2 (1-p)^{k-1} p \\ &= \sum_{k=1}^{\infty} (k-1)^2 (1-p)^{k-1} p + \sum_{k=1}^{\infty} 2(k-1)(1-p)^{k-1} p + \sum_{k=1}^{\infty} (1-p)^{k-1} p \\ (\text{let } n = k-1) &= (1-p) \sum_{n=1}^{\infty} n^2 (1-p)^{n-1} p + 2(1-p) \sum_{k=1}^{\infty} n(1-p)^{n-1} p + 1 \\ &= (1-p) E[X^2] + 2(1-p) E[X] + 1 \end{split}$$

Then substitute $E[X] = \frac{1}{p}$ and solve the equation to get $E[X^2] = \frac{2-p}{p^2}$.

Hence
$$Var(X) = E[X^2] - (E[X])^2 = \frac{1-p}{p^2}$$

Remark. Observe that we have use two different ways to compute E[X] and Var(X) in Example 1 both of which can be recursively extended to deal with series like $\sum_{k=1}^{\infty} k^p x^k$ with $p \in \mathbb{N}$.

Example 2. Let $X \sim Bin(n, p)$. Prove

$$E\left[\frac{1}{1+X}\right] = \frac{1 - (1-p)^{n+1}}{(n+1)p}$$

Proof. Recall the PMF of Bin(n, p) and by the formula of E[g(X)],

$$\begin{split} E\left[\frac{1}{1+X}\right] &= \sum_{k=0}^{n} \frac{1}{k+1} \binom{n}{k} p^{k} (1-p)^{n-k} \\ &= \frac{1}{(n+1)p} \sum_{k=0}^{n} \frac{1}{k+1} \frac{(n+1)!}{k![(n+1)-(k+1)]!} \cdot p^{k+1} (1-p)^{[(n+1)-(k+1)]} \\ &= \frac{1}{(n+1)p} \sum_{k=0}^{n} \binom{n+1}{k+1} p^{k+1} (1-p)^{[(n+1)-(k+1)]} \\ &(\text{let } j = k+1) \quad = \frac{1}{(n+1)p} \left[\sum_{j=0}^{n} \binom{n+1}{j} p^{j} (1-p)^{[(n+1)-j]} - (1-p)^{n+1} \right] \\ (\text{ by Binomial Thm}) \quad = \frac{1}{(n+1)p} [1 - (1-p)^{n+1}]. \end{split}$$

Example 3. Let X be a r.v. with non-negative integral values. Prove

$$\sum_{k=0}^{\infty} k P(X \ge k) = \frac{1}{2} (E[X^2] + E[X]).$$

Proof.

$$\sum_{k=0}^{\infty} kP(X \ge k) = \sum_{k=0}^{\infty} \sum_{i=k}^{\infty} kP(X = i)$$
$$= \sum_{i=0}^{\infty} \sum_{k=0}^{i} kP(X = i)$$
$$= \sum_{i=0}^{\infty} \frac{i(i+1)}{2} P(X = i)$$
$$= \frac{1}{2} E[X^{2}] + \frac{1}{2} E[X],$$

where in the second equality we have changed the order of summation.