Solution to Midterm

1. (a) Let E_1, E_2 be the events that the outcome is 6 and 9 respectively. Let E be the event that E_1 occurs first. By direct counting,

$$P(E_1) = \frac{5}{36}, \quad P(E_2) = \frac{4}{36}.$$

Method 1:

$$P(E) = P(E|E_1)P(E_1) + P(E|E_2)P(E_2) + P(E|(E_1 \cup E_2)^c)P((E_1 \cup E_2)^c)$$

= $P(E_1) + P(E)P((E_1 \cup E_2)^c)$ (independence)
= $\frac{P(E_1)}{P(E_1 \cup E_2)}$

Method 2:

$$P(E) = \sum_{i=0}^{\infty} (1 - P(E_1) - P(E_2))^i P(E_1)$$
$$= \frac{P(E_1)}{P(E_1) + P(E_2)}$$

(b) Statement: Let E_1 and E_2 be 2 mutually exclusive events of a trial which is to be repeated indefinitely. Then the probability of E_2 occurs before E_1 is $\frac{P(E_2)}{P(E_1 \cup E_2)}$. The proof is identical to that part (a).

2.
$$(a)$$

$$P(\text{three of a kind}) = \frac{C_3^{13} C_1^3 C_3^4 C_1^4 C_1^4}{C_5^{52}}, \quad P(\text{four of a kind}) = \frac{C_2^{13} C_1^2 C_4^4 C_1^4}{C_5^{52}}$$

(b) Let E_i be the event that all 4 of the number *i* cards are in the hand. For distinct i, j, k,

$$P(E_i) = \frac{C_9^{48}}{C_{13}^{52}}, \quad P(E_i E_j) = \frac{C_5^{44}}{C_{13}^{52}}, \quad P(E_i E_j E_k) = \frac{C_1^{40}}{C_{13}^{52}}.$$

Note also any 4 or more intersections of these events has probability 0. By the inclusion-exclusion principle, the required probability is

$$P\left(\bigcup_{i=1}^{13} E_i\right) = \sum_{i=1}^{13} P(E_i) - \sum_{i < j} P(E_i E_j) + \sum_{i < j < k} P(E_i E_j E_k)$$
$$= C_1^{13} \frac{C_9^{48}}{C_{13}^{52}} - C_2^{13} \frac{C_5^{44}}{C_{13}^{52}} + C_3^{13} \frac{C_1^{40}}{C_{13}^{52}}.$$

3. (a)

$$\begin{split} P(E|F) &= \frac{P(EF)}{P(F)} \\ &= \frac{P(F|E)P(E)}{P(F|E)P(E) + P(F|E^c)P(E^c)} \end{split}$$

(b) Let E be the event that the first box is chosen and F be the event that the chosen marble is white. By Bayes formula,

$$P(E|F) = \frac{P(F|E)P(E)}{P(F|E)P(E) + P(F|E^c)P(E^c)}$$
$$= \frac{\frac{1}{2} \cdot \frac{1}{2}}{\frac{1}{2} \cdot \frac{1}{2} + \frac{1}{3} \cdot \frac{1}{2}}$$
$$= \frac{3}{5}$$

$$P(X=n) = \frac{1}{2} \frac{2}{3} \cdots \frac{n-1}{n} \frac{1}{n+1} = \frac{1}{n(n+1)}$$

(b)

$$P(X < \infty) = 1 - P(X = \infty) = 1 - \lim_{n \to \infty} \frac{1}{2} \frac{2}{3} \cdots \frac{n}{n+1} = 1 - \lim_{n \to \infty} \frac{1}{n+1} = 1$$

(c)

$$E[X] = \sum_{k=1}^{\infty} kP(X = k)$$
$$= \sum_{k=1}^{\infty} k \frac{1}{k(k+1)}$$
$$= \sum_{k=2}^{\infty} \frac{1}{k} = \infty$$

5. (a)

$$P(X = k) = \frac{\lambda^k e^{-\lambda}}{k!}, k = 0, 1, 2, \cdots, \quad E[X] = \lambda$$

(b)

$$E[X^{n+1}] = \sum_{k=0}^{\infty} k^{n+1} \frac{\lambda^k}{k!}$$
$$= \lambda \sum_{k=1}^{\infty} k^n \frac{\lambda^{k-1}}{(k-1)!}$$
$$= \lambda \sum_{k=0}^{\infty} (k+1)^n \frac{\lambda^k}{k!}$$
$$= \lambda E[(X+1)^n]$$

(c) By (b), $E[X^3] = \lambda E[(X+1)^2] = \lambda E[X^2+2X+1] = \lambda (\text{Var}(X) + E[X]^2 + 2E[X] + 1) = \lambda (\lambda^2 + 3\lambda + 1).$